

# Reminder of Random Variables

## Indicator RV

Dr Ahmad Khonsari

ECE Dept.

The University of Tehran

## 7. Indicator RVs

The **indicator random variable**  $I\{A\}$  associated with **event**  $A$  is defined as

$$I\{A\} = \begin{cases} 1 & \text{if } A \text{ occurs} \\ 0 & \text{if } A \text{ does not occur} \end{cases} \quad (7.1)$$

**Example:** determine the expected number of heads in tossing a fair coin. **sample space** is

$$S = \{H, T\}, \text{ with } \Pr\{T\} = \Pr\{H\} = \frac{1}{2}.$$

Define the event  $H$  as the coin coming up heads, we define an indicator RV  $X_H$  associated with the **event**  $H$ , such that:

$X_H$  counts the number of heads obtained in this flip, i.e. it is 1 if the coin comes up heads and 0, otherwise.

We write

$$X_H = I\{H\} = \begin{cases} 1 & \text{if } H \text{ occurs} \\ 0 & \text{if } T \text{ occurs} \end{cases}.$$

# 7. Indicator RVs

The **expected number** of heads obtained in **one flip** of the coin is simply the expected value of indicator variable  $X_H$ :

$$\begin{aligned} E[X_H] &= E[I\{H\}] \\ &= 1 \cdot \Pr\{H\} + 0 \cdot \Pr\{T\} \\ &= 1 \cdot \left(\frac{1}{2}\right) + 0 \cdot \left(\frac{1}{2}\right) = \frac{1}{2} \end{aligned}$$

Thus the **expected number of heads** obtained by one flip of a fair coin is  $1/2$ .

Q: what is the difference between expected value and average case?  
Does make sense to define average with one flip ?

## 7. Indicator RVs

### Lemma 7.1

Given a sample space  $S$  and an event  $A$  in the sample space  $S$ , let

$$X_A = I\{A\}.$$

Then 
$$E[X_A] = Pr\{A\}$$

### Proof:

By the definition of an indicator RV from equation (7.1) and the definition of expected value, we have

$$\begin{aligned} E[X_A] &= E[I\{A\}] \\ &= 1 \cdot Pr\{A\} + 0 \cdot Pr\{\bar{A}\} \\ &= Pr\{A\} \end{aligned}$$

, where  $\bar{A}$  denotes  $S - A$ , (i.e. the complement of  $A$ ).

Thus the above lemma implies:

**The expected value of an indicator RV associated with an event  $A$  is equal to the probability that  $A$  occurs.**

## 7. Indicator RVs

Although indicator RVs may seem cumbersome for an application such as counting the expected number of heads on a flip of a single coin, they are useful for analyzing situations in which we perform repeated random trials.

**Example:** compute the expected number of heads in  $n$  tossing of a coin.

Let  $X$  denotes the total number of heads in the  $n$  coin flips, so that

$$X = \sum_{i=1}^n X_i$$

we take the expectation of both sides

$$\begin{aligned} E[X] &= E \left[ \sum_{i=1}^n X_i \right] \\ &= \sum_{i=1}^n E[X_i] \\ &= \sum_{i=1}^n \frac{1}{2} \\ &= \frac{n}{2} \end{aligned}$$

## 7. Indicator RVs

We can compute the expectation of a random variable having a binomial distribution from equations

$$C_n^k = \binom{n}{k} = \frac{n}{k} \binom{n-1}{k-1}$$

And

$$f_{n,p}(k) = C_n^k p^k (1-p)^{n-k}$$

$$\sum_{k=0}^n f_{n,p}(k) = 1.$$

## 7. Indicator RVs

Let  $X \sim \text{Bin}(n; p)$ ,  $q=1-p$ , By the definition of expectation, we have

$$\begin{aligned} E[X] &= \sum_{k=0}^n k \cdot \Pr\{X = k\} \\ &= \sum_{k=0}^n k \cdot f_{n,p}(k) \\ &= \sum_{k=0}^n k \binom{n}{k} p^k q^{n-k} = \sum_{k=1}^n k \frac{n}{k} \binom{n-1}{k-1} p^k q^{n-k} \\ &= np \sum_{k=1}^n \binom{n-1}{k-1} p^{k-1} q^{n-k} && \begin{aligned} k-1 &= j && = k \\ k &= j+1 \\ n-k &= n-(j+1) \\ &= n-1-j \end{aligned} \\ &= np \sum_{k=0}^{n-1} \binom{n-1}{k} p^k q^{(n-1)-k} \\ &= np \sum_{k=0}^{n-1} f_{n-1,p}(k) \\ &= np \end{aligned}$$

## 7. Indicator RVs

Let  $X \sim \text{Bin}(n; p)$ ,  $q=1-p$ . Obtaining the same result using the linearity of expectation.

Let  $X_i$  denotes the number of successes in the  $i$  th trial. Then

$$E[X_i] = p \cdot 1 + q \cdot 0 = p$$

and by linearity of expectation, the expected number of successes for  $n$  trials is

$$\begin{aligned} E[X] &= E \left[ \sum_{i=1}^n X_i \right] \\ &= \sum_{i=1}^n E[X_i] \\ &= \sum_{i=1}^n p \\ &= np \end{aligned}$$



## 7. Indicator RVs

**Example:** Let  $X \sim \text{Bin}(n; p)$ ,  $q=1-p$  calculate the variance of the distribution. Using  $\text{Var}[X] = E[X^2] - E^2[X]$ ,

we have  $\text{Var}[X_i] = E[X_i^2] - E^2[X_i]$ .

$X_i$  only takes on the values 0 and 1, we have  $X_i^2 = X_i$ ,

which implies  $E[X_i^2] = E[X_i] = p$ .

Hence,  $\text{Var}[X_i] = p - p^2 = pq$

To compute the variance of  $X$ , we take advantage of the **independence** of the  $n$  trials; thus,

$$\begin{aligned}\text{Var}[X] &= \text{Var}\left[\sum_{i=1}^n X_i\right] \\ &= \sum_{i=1}^n \text{Var}[X_i] \\ &= \sum_{i=1}^n pq \\ &= npq\end{aligned}$$

تفاوت  $\mathcal{F}$  شمارد بر روی  $\Omega$   $\rightarrow$  همگام مستقیم با  $\mathcal{F}$  است و  $\mathcal{F}$  هم از  $\mathcal{F}$  ریشه ندارد

$\mathcal{F}$  شمار یک تقابل کننده و  $\mathcal{F}$  هم از  $\mathcal{F}$  ریشه ندارد

متغیر تصادفی (Indicator RV) : کمیت  $A$  ، تقابل کننده را به  $\mathcal{F}$  شمار محدود

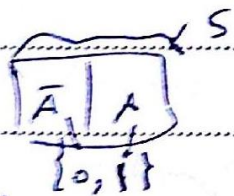
و نفع آید و  $\mathcal{F}$  (collectively exhaustive)  $\mathcal{F}$  است (همه  $\mathcal{F}$  ها را در نظر بگیرید)

$\mathcal{F}$  شمار  $A$  را با متغیر تصادفی  $I_A$   $\mathcal{F}$  شمار هم  $\mathcal{F}$  است و  $\mathcal{F}$  شمار  $A$  را

$$I_A(\omega) = \begin{cases} 1 & \text{if } \omega \in A \\ 0 & \text{if } \omega \in \bar{A} \end{cases}$$

تفاوت مستقیم و  $\mathcal{F}$  شمار ، در هر دو یک دست به هم وصل می شود و هر دو دارد ،  $\mathcal{F}$  شمار  $A$  را  
اگر  $\mathcal{F}$  شمار  $A$  را  $\mathcal{F}$  شمار  $A$  را  $\mathcal{F}$  شمار  $A$  را

بیت به A فقط در صورتی که رخ دهد  $I_A = 1$  (بیشتر  $\frac{1-1}{n}$  تا پیش از آن زمان رخ دهد)

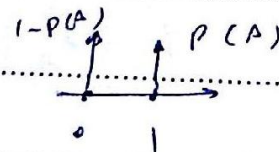


$P_{I_A}$  (تابع جرم احتمال)  $I_A$  (فرض کنید که در این حالت)  $\{0, 1\}$  (بهترین)

$$P_{I_A}(0) = P(\bar{A}) = 1 - P(A)$$

تاریخ سال

$$P_{I_A}(1) = P(A)$$



# Appendix

## 8. Linearity of Expectation

# Linearity of Expectation

- Thm. 
$$E(X_1 + X_2) = \sum_{s \in S} p(s)(X_1(s) + X_2(s))$$

$$E(X_1 + X_2 + \dots + X_n) = E(X_1) + E(X_2) + \dots + E(X_n)$$

Very useful result. Holds even when  $X_i$ 's are dependent!

Proof: (case  $n=2$ )

Proof: (case n=2)

$$\begin{aligned} E(X_1 + X_2) &= \sum_{s \in \mathcal{S}} p(s)(X_1(s) + X_2(s)) \\ &= \sum_{s \in \mathcal{S}} p(s)X_1(s) + \sum_{s \in \mathcal{S}} p(s)X_2(s) \\ &= E(X_1) + E(X_2) \end{aligned}$$

QED

**(defn. of  
expectation;  
summing over  
elementary  
events in  
sample space)**

Aside, alternative defn. of expectation:

$$E(X) = \sum_x xP(X = x)$$

# Example

- Consider  $n$  children of different heights placed in a line at random.
- Starting from the beginning of the line, select the first child. Continue walking, until you find a taller child or reach the end of the line.
- When you encounter taller child, also select him/her, and continue to look for next tallest child or reach end of line.
- **Question: What is the expected value of the number of children selected from the line??**

Hmm. Looks tricky...

*What would you guess? Lineup of 100 kids... [e.g. 15 more or less?]  
Lineup of 1,000 kids... [e.g. 25 more or less?]*

- Let  $X$  be the r.v. denoting the number of children selected from the line.

$$X = X_1 + X_2 + \dots + X_n$$

where

$$X_i = \begin{cases} 1 & \text{if the tallest among the first } i \text{ children.} \\ & \text{(i.e. will be selected from the line)} \\ 0 & \text{otherwise} \end{cases}$$

By linearity of expectation, we have:

$$E(X) = E(X_1 + X_2 + \dots + X_n) = E(X_1) + E(X_2) + \dots + E(X_n)$$



- What is  $E(X_i)$ ?

What is  $P(X_1 = 1)$ ?      A: 1

What is  $P(X_n = 1)$ ?      A:  $1/n$

What is  $P(X_i = 1)$ ?      A:  $1/i$

Note that the one tallest person among  $i$  persons needs to be at the end. (by symmetry: equally likely in any position)

$$\begin{aligned} \text{Now, } E(X_i) &= 0 * P(X_i = 0) + 1 * P(X_i = 1) \\ &= 0 * (1 - 1/i) + 1 * (1/i) \\ &= 1/i. \end{aligned}$$

So,  $E(X) = 1 + 1/2 + 1/3 + 1/4 \dots + 1/n$

Consider doubling queue:  
What's probability tallest kid in first half

A:  $1/2$   
(in which case 2<sup>nd</sup> half doesn't add anything!)

$\approx \ln(n) + \gamma + \frac{1}{n}$ , where  $\gamma = 0.57722\dots$  Euler's constant

**Intriguing!**

**Surprisingly small!!** e.g. **N = 100**      **E(X) ~ 5**

**Intuition??**      **N = 200**      **E(X) ~ 6**

**N = 1000**      **E(X) ~ 7**

**N = 1,000,000**      **E(X) ~ 14**

**N = 1,000,000,000**      **E(X) ~ 21**

# Indicator Random Variable

- Recall: Linearity of Expectation

$$E(X_1 + X_2 + \dots + X_n) = E(X_1) + E(X_2) + \dots + E(X_n)$$

$Y \rightarrow$  An indicator random variable is:

- 0/1 binary random variable.
- 1 corresponds to an event  $E$ , and 0 corresponds to the event did not occur
- Then  $E(Y) = 1 \cdot P(E) + 0 \cdot (1 - P(E)) = P(E)$

Expected number of times event occurs:

- $E(X_1 + X_2 + \dots) = E(X_1) + E(X_2) + \dots = P(E_1) + P(E_2) + \dots$

$$E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$$

Suppose everyone ( $n$ ) puts their cell phone in a pile in the middle of the room, and I return them randomly. What is the expected number of students who receive their own phone back? Guess??

Define for  $i = 1, \dots, n$ , a random variable:

$$X_i = \begin{cases} 1 & \text{if student } i \text{ gets the right phone,} \\ 0 & \text{otherwise.} \end{cases}$$

Need to calculate:

$$E\left[\sum_{i=1}^n X_i\right]$$

k	0	1
$\Pr(X_i=k)$	$1-(1/n)$	$1/n$

$$E[X_i] = \Pr(X_i = 1)$$

19 Why? Symmetry! All phones equally likely.

$$E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$$

So,

$$E[X] = E[X_1 + X_2 + \dots + X_n]$$

$$= E[X_1] + E[X_2] + \dots + E[X_n]$$

$$= 1/n + 1/n + \dots + 1/n$$

$$= 1$$

So, we expect just **one** student to get his or her own cell phone back...

Independent of  $n$ !!

$$E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$$

Suppose there are  $N$  couples at a party, and suppose  $m$  (random) people get "sleepy and leave"... What is the expected number of ("complete") couples left?

Define for  $i = 1, \dots, N$ , a random variable:

$$X_i = \begin{cases} 1 & \text{if couple } i \text{ remains,} \\ 0 & \text{otherwise.} \end{cases}$$

Define r.v.  $X = X_1 + X_2 + \dots + X_n$ , and we want  $E[X]$ .

$$E[X] = E[X_1 + X_2 + \dots + X_n] \\ = E[X_1] + E[X_2] + \dots + E[X_n]$$

So, what do we know about  $X_i$ ?

$$E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$$

Suppose there are  $N$  couples at a party, and suppose  $m$  people get sleepy and leave. What is the expected number of couples left?

Define for  $i = 1, \dots, N$ , a random variable:

$$X_i = \begin{cases} 1 & \text{if couple } i \text{ remains,} \\ 0 & \text{otherwise.} \end{cases}$$

$$E[X_i] = \Pr(X_i = 1)$$

How?

(# of ways of choosing  $m$  from everyone else) / (# of ways of choosing  $m$  from all)

$$= \frac{\binom{2N-2}{m}}{\binom{2N}{m}}$$

$$E[X_1] + E[X_2] + \dots + E[X_n]$$

$$= n \times E[X_1] = (2N-m)(2N-m-1)/2(2N-1)$$

#### 4. [10 points] Expectation

Imagine there are  $N$  couples at a party, and suppose  $m$  people get sleepy and have to go home. When a person goes home, his or her date has to leave with them. We need to find the expected number of couples left at the party.

(a) First, let's consider a single couple, couple  $i$ . Define a random variable  $X_i$  to be:

$$X_i = \begin{cases} 1 & \text{if couple } i \text{ stays} \\ 0 & \text{if couple } i \text{ leaves} \end{cases}$$

This sort of function is known as an *indicator*.

What is the expected value of  $X_i$ ? Clearly explain your answer.

**Solution 1:** The couple stays only if neither gets sleepy. To compute this probability, we count the number of ways to choose  $m$  sleepy people from everyone else excluding this couple, then divide by the total number of ways to choose  $m$  sleepy people. So we get:

$$p = \frac{\binom{2N-2}{m}}{\binom{2N}{m}} = \frac{\frac{(2N-2)!}{(2N-2-m)!m!}}{\frac{(2N)!}{(2N-m)!m!}} = \frac{(2N-m)(2N-m-1)}{2N(2N-1)}$$

The expected value of  $X_i$  is thus  $\mathbf{E}(X_i) = p * 1 + (1 - p) * 0 = p$ .

**Solution 2:** The couple stays only if neither gets sleepy, which occurs with probability  $p = (1 - \frac{m}{2N})(1 - \frac{m}{2N-1})$ . This is because  $\frac{m}{2N}$  is the probability that the first person in the couple gets sleepy, and  $1 - \frac{m}{2N}$  is the probability that the first person in the couple does not get sleepy and stays.

$1 - \frac{m}{2N-1}$  is the probability that the *second* person in the couple does not get sleepy and stays, given that the first person does not get sleepy and stays. The denominator is only  $2N - 1$  since one person is already known to not be sleepy. If you fiddle with the algebra, you'll see that this equation is equal to the one in the previous solution.

Again, the expected value of  $X_i$  is thus  $\mathbf{E}(X_i) = p * 1 + (1 - p) * 0 = p$ .

- (b) What is the expected number of couples left at the party? Your answer should be a function of the variables  $N$  and  $m$ .

*Hint:* Suppose we have  $n$  random variables  $X_i$  and we wish to compute the expected value of  $\sum_{i=1}^n X_i$ . This means we have sum of  $n$  functions and we wish to find the expected value of the sum. We can do this by using the *linearity of expectation* which establishes that  $\mathbf{E}(\sum_{i=1}^n X_i) = \sum_{i=1}^n \mathbf{E}(X_i)$ . This means that we can find the answer by computing the expectation for each  $X_i$  separately and then summing the results.

**Solution 1:** We use our result for **Solution 1** of part a. By linearity of expectation, the expected number of couples left at the party is  $\sum_{i=1}^N \mathbf{E}(X_i) = \sum_{i=1}^N p = N * p$  which is equal to:

$$\frac{(2N - m)(2N - m - 1)}{2(2N - 1)}$$



**Solution 2:** We use our result for **Solution 2** of part a. By linearity of expectation, the expected number of couples left at the party is  $\sum_{i=1}^N \mathbf{E}(X_i) = \sum_{i=1}^N p = N * p$  which is equal to:

$$N \left(1 - \frac{m}{2N}\right) \left(1 - \frac{m}{2N-1}\right)$$

**Note:** Algebraic manipulation shows that the two solutions are equivalent.